A new methodology to test galaxy formation models using the dependence of clustering on stellar mass

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ABSTRACT

We present predictions for the two-point correlation function of galaxy clustering as a function of stellar mass, computed using two new versions of the GALFORM semi-analytic galaxy formation model. These models make use of a high resolution, large volume N-body simulation, set in the 7-year Wilkinson Microwave Anisotropy Probe cosmology. One model uses a universal stellar initial mass function (IMF), while the other assumes different IMFs for quiescent star formation and bursts. Particular consideration is given to how the assumptions required to estimate the stellar masses of observed galaxies (such as the choice of IMF, stellar population synthesis model, and dust extinction) influence the perceived dependence of galaxy clustering on stellar mass. Broad-band spectral energy distribution fitting is carried out to estimate stellar masses for the model galaxies in the same manner as in observational studies. We show clear differences between the clustering signals computed using the true and estimated model stellar masses. As such, we highlight the importance of applying our methodology to compare theoretical models to observations. We introduce an alternative scheme for the calculation of the merger time-scales for satellite galaxies in GALFORM, which takes into account the dark matter subhalo information from the simulation. This reduces the amplitude of small-scale clustering. The new merger scheme offers improved or similar agreement with observational clustering measurements, over the redshift range 0 < z < 0.7. We find reasonable agreement with clustering measurements from the Galaxy and Mass Assembly Survey, but find larger discrepancies for some stellar mass ranges and separation scales with respect to measurements from the Sloan Digital Sky Survey and the VIMOS Public Extragalactic Redshift Survey, depending on the GALFORM model used.

Key words: galaxies: evolution – galaxies: formation – galaxies: stellar content – large-scale structure of Universe.

1 INTRODUCTION

Galaxy formation involves the interplay between a variety of physical processes, such as the suppression of star formation in massive haloes due to the shutting down of gas cooling by active galactic nuclei (AGN), and the heating of the interstellar medium by supernovae. These processes vary in importance with redshift and the mass of the dark matter halo which hosts the galaxy. Semi-analytic models use simplified expressions and approximations to describe these complex physical processes, allowing predictions to be generated for how dark matter haloes are populated with galaxies (Benson et al. 2000a; Cole et al. 2000; Hatton et al. 2003; Bower et al. 2006; Croton et al. 2006; Monaco, Fontanot & Taffoni 2007; Somerville et al. 2008).

The model processes are constrained through comparison to observed statistics of the galaxy population, such as the luminosity function (e.g. Gonzalez-Perez et al. 2014). We can also constrain the galaxy formation physics by measuring galaxy clustering as a function of intrinsic properties, such as stellar mass. Appealing to the clustering as a function of galaxy properties provides a
constraint on how dark matter haloes are populated with galaxies (Kauffmann et al. 1999; Benson et al. 2000b). Galaxy formation models have been used to make predictions for comparison to observational measurements of the clustering as a function of different properties, such as luminosity, star formation rate, and stellar mass (e.g. Norberg et al. 2001; Kim et al. 2009; Meneux et al. 2009; Guo et al. 2011, 2013; Li et al. 2012; Henriques et al. 2013; Marulli et al. 2013; Briely et al. 2014). Some models explicitly consider galaxy clustering when setting their parameters (e.g. Guo et al. 2011). For models which have been constructed without using galaxy clustering as a constraint, comparison to clustering measurements is a test of their predictive power. The models considered in this paper have not been calibrated to reproduce any clustering measurements, or stellar mass function data.

Despite the perceived complexity of semi-analytic models, Contreras et al. (2013) show that the clustering predictions made by different groups are robust, at least for samples defined by stellar mass. They considered the Durham GALFORM (Bower et al. 2006; Font et al. 2008) and Munich LGALAXIES (Bertone, De Lucia & Thomas 2007; de Lucia & Blaizot 2007; Guo et al. 2011) model families. These two sets of models differ in their implementation of various physical processes of galaxy formation. Contreras et al. find differences in the clustering predictions on small scales, which they trace to the handling of galaxy mergers. We explore the impact of the choice of galaxy merger scheme further in this paper.

The stellar masses of real galaxies have to be derived from observable properties such as broad-band photometry, which requires the assumption of a star formation history (SFH), stellar initial mass function (IMF), stellar population synthesis (SPS) model, and dust extinction model (e.g. Pforr, Maraston & Tonini 2012; Mitchell et al. 2013). The choice of the particular set of assumptions used has an impact on the recovered masses, and so there are important systematic uncertainties inherent in the derived stellar masses of observed galaxies. These uncertainties may influence comparisons of galaxy clustering as a function of stellar mass, between theoretical models and observational measurements. Thus if the clustering as a function of stellar mass is to be used to constrain theoretical models, we must take care to treat appropriately the difficulties in estimating the stellar masses of observed galaxies.

Contreras et al. (2013) compared theoretical predictions by examining galaxy samples of fixed abundance. We extend this study, motivated by the work of Mitchell et al. (2013, see also Marchesini et al. 2009; Pforr et al. 2012), and demonstrate a new methodology which we find is essential for comparing model predictions to observational clustering data. Our technique consists of carrying out broad-band spectral energy distribution (SED) fitting to compute stellar mass estimates for model galaxies, in the manner that is typically applied to observed galaxies. Such a treatment is particularly important for understanding the influence of dust extinction, and the impact of Eddington bias in higher mass bins where the stellar mass function is steep.

We present the clustering predictions of two new GALFORM variants: the models of Gonzalez-Perez et al. (2014) and Lacey et al. (in preparation), hereafter Gon14 and Lac14. These models have been calibrated to reproduce observations in the 7-year Wilkinson Microwave Anisotropy Probe (WMAP7) cosmology (Komatsu et al. 2011). The models make use of the MS-W7 N-body simulation (Guo et al. 2013), which is a new version of the Millennium Simulation (Springel et al. 2005) updated to use the WMAP7 cosmological parameters. These models take advantage of extensions to the galaxy formation physics implemented in GALFORM, making use of an empirical law to determine star formation rates (Lagos et al. 2011). For discussions of the implications for galaxy formation models of the WMAP7 cosmology, see Guo et al. (2013) and Gon14.

We compare the model predictions to derived results from the Sloan Digital Sky Survey (SDSS; York et al. 2000; Abazajian et al. 2009), the Galaxy and Mass Assembly Survey (GAMA; Driver et al. 2011), and the VIMOS Public Extragalactic Redshift Survey (VIPERS; Guzzo et al. 2014). Together, these surveys have measured the clustering of galaxies as a function of stellar mass up to redshift \(z \sim 1\) (Li et al. 2006; Marulli et al. 2013; Farrow et al., in preparation). In particular, we compare to these observational results at redshifts 0.1 (SDSS), 0.2 (GAMA), and 0.6 (VIPERS). The specific star formation rate falls dramatically over this interval (e.g. Weinmann, Neistein & Dekel 2011), so the interplay between the different galaxy formation processes could change over the cosmic times considered. Through comparing to these measurements, we assess both the importance of carrying out SED fitting to the model photometry, and the level of agreement between the measurements and the model predictions.

The layout of this paper is as follows. Details of the galaxy formation models and observational data are given in Sections 2 and 3. Our methodology is described in Section 4. Our results are presented and discussed in Section 5. Concluding remarks are given in Section 6.

All magnitudes are on the AB system (Oke 1974). The assumed Λ cold dark matter (ΛCDM) cosmological parameters are listed in Section 2.2. Comoving length units are used throughout this paper.

## 2 THEORETICAL MODELLING

This section describes the galaxy formation models used in our study of galaxy clustering as a function of stellar mass. We first give an overview of GALFORM (Section 2.1), followed by a description of the N-body simulation used (Section 2.2), then contrast the two models compared (Section 2.3), and describe a new scheme for the treatment of satellite galaxy orbits and merger time-scales (Section 2.4).

### 2.1 Overview of GALFORM

Galaxy formation within dark matter haloes, as followed in recent GALFORM variants, can be broken down into several key processes: (i) formation and hierarchical growth of dark matter haloes; (ii) shock heating of baryonic material falling into haloes, followed by radiative cooling and disc formation; (iii) quiescent star formation in discs (and bursts due to instabilities); (iv) suppression of gas cooling (and hence of star formation) through feedback from supernovae, AGN, and photonization of the intergalactic medium; (v) chemical enrichment of the stars and gas; and (vi) mergers of satellite galaxies with the central galaxy of their halo, due to dynamical friction, which can cause bursts of star formation. For an overview of the development of the GALFORM model, see Benson & Bower (2010). For details of the galaxy formation physics, see Baugh (2006) and Benson (2010).

In order to connect the predictions of galaxy formation models to the properties of observed galaxies, an SPS model must be assumed, along with a model to describe dust attenuation. SPS models, such as those of Bruzual & Charlot (2003), Maraston (2005), and Conroy, Gunn & White (2009), compute the SED of a coeval stellar population with a given initial metallicity, as a function of age. Convolving this stellar SED with the SFH of a galaxy (i.e. its star formation rate as a function of time), taking into account its chemical enrichment history (the metallicity of stars forming at a given time), yields the
SED of the galaxy itself (Cole et al. 2000). SPS models require a stellar IMF to be specified, which gives the distribution of masses of stars formed in a given episode of star formation.

Attenuation of starlight by dust in GALFORM is modelled in a physically motivated way, in which the stars and dust are mixed together, based on radiative transfer and the geometry of a disc and bulge (Silva et al. 1998; Ferrara et al. 1999; Cole et al. 2000; Lacey et al. 2011; Gonzalez-Perez et al. 2013). For a given photometric band, the galactic SED is weighted by the wavelength response of the filter and integrated to yield the flux.

2.2 The MS-W7 simulation

The cosmological parameters from WMAP (Komatsu et al. 2011) have been used in an N-body simulation similar to the Millennium Simulation of Springel et al. (2005). This MS-W7 simulation (Guo et al. 2013) has present-day density parameters of $\Omega_M = 0.272$, $\Omega_b = 0.0455$, and $\Omega_{\Lambda} = 0.728$, for matter, baryons, and dark energy, respectively. The present-day Hubble parameter is $H = 100h \text{ km s}^{-1} \text{Mpc}^{-1}$, where $h = 0.704$. The spectral index of primordial fluctuations is $n_s = 0.967$, and the linear perturbation amplitude is $\sigma_8 = 0.810$. The simulation follows $2160^3$ particles from redshift 127, in a volume of comoving side $L_{\text{box}} = 500 h^{-1}$ Mpc.

2.3 The Gon14 and Lac14 models

The Gon14 and Lac14 GALFORM models are based on the halo merger trees of the MS-W7 simulation. They are separate developments of the Lagos et al. (2012) model, which used an empirical star formation rate law introduced by Lagos et al. (2011). The Lagos et al. (2012) model in turn is based on that of Bower et al. (2006), which introduced AGN feedback into GALFORM. The new models will be made publicly available in the Millennium Database.\footnote{http://virgodb.cosma.dur.ac.uk:8080/Millennium}

The IMF $\zeta(M_*)$ is defined such that the number of newly formed stars, per solar mass, with stellar mass, $M_*$, between $\log_{10}(M_*/M_{\odot})$ and $\log_{10}(M_*/M_{\odot}) + d\log_{10}(M_*/M_{\odot})$ is given by $\zeta(M_*) d\log_{10}(M_*/M_{\odot})$, for some mass unit $M_{\odot}$. A power law is often used, such that

$$\zeta(M_*) d\log_{10}\left(\frac{M_*}{M_{\odot}}\right) = \zeta_0 \left(\frac{M_*}{M_{\odot}}\right)^{-x} d\log_{10}\left(\frac{M_*}{M_{\odot}}\right), \quad (1)$$

for some normalization $\zeta_0$ and slope $x$. The Gon14 model uses a Kennicutt (1983) IMF, which is a broken power law. This has a slope of $x = 0.4$ for $M_* < 1 \text{ M}_{\odot}$, and $x = 1.5$ for $M_* > 1 \text{ M}_{\odot}$. In the Lac14 model, distinct IMFs are employed in quiescent star formation and bursts. In the former case, a Kennicutt IMF is used as in Gon14. However, for bursts, the IMF is taken to be a single power law with $x = 1$. Such non-universality of the IMF is argued to be necessary to match the observed number counts and redshift distribution of submillimetre galaxies (Baugh et al. 2005).

The SPS model used in the Gon14 model is a private release of the Bruzual & Charlot series from 1999, which is intermediate between Bruzual & Charlot (1993) and Bruzual & Charlot (2003). The Lac14 model uses the Maraston (2005) SPS model. The Maraston model attributes much more luminosity to stars in the thermally pulsating asymptotic giant branch (TP-AGB) phase than is done in the Bruzual & Charlot model. Such stars emit strongly in the near-infrared (near-IR; e.g. MacArthur et al. 2010). The influence of TP-AGB stars in the Maraston model has been the source of some debate in the literature (e.g. Marigo & Girardi 2007; Kriek et al. 2010; MacArthur et al. 2010; Zibetti et al. 2013). Gon14 study the influence of using various alternative SPS models, including that of Maraston. They find that the choice of SPS model does not affect the evolution of the rest-frame optical and ultraviolet (UV) luminosity functions, but models incorporating strong TP-AGB emission yield significantly different evolution of the rest-frame near-IR luminosity function (see also Tonini et al. 2009; Henriques et al. 2011). This choice gives an improved match to the bright end of the rest-frame K-band luminosity function at high redshifts in the Lac14 model.

Additional parameter differences between the Gon14 and Lac14 models are described in Appendix A.

2.4 Subhalo dynamical friction for satellites

In the standard models considered here, time-scales for the merging of satellites with the central galaxy in their host halo due to dynamical friction are computed in the models as described by Cole et al. (2000). This method assumes that when a new halo forms, each satellite galaxy enters the halo on a random orbit. The merger time-scale is then computed using an analytical formula which assumes the halo to be isothermal. While the Gon14 model makes use of the equations presented by Lacey & Cole (1993), a modified expression is used in the Lac14 model. This expression has been empirically fitted to numerical simulations to account for the tidal stripping of subhaloes (Jiang et al. 2008; Jiang, Jing & Lin 2010), but otherwise the treatment is the same, i.e. an analytic time-scale is computed as soon as a galaxy enters a larger halo. The satellite is considered to have merged with its central galaxy once the merger time-scale has elapsed, provided that this transpires before the halo merges to form a larger system, in which case a new merger time-scale is computed. Note that this scheme does not take into account that the satellite may still reside in a resolvable dark matter subhalo at the time the merger takes place.

We have implemented an alternative treatment of mergers, which makes use of the subhalo information from the simulation. The motivation for using this new scheme is that it is more faithful to the underlying N-body simulation, minimizing the reliance on analytically determined orbits. Instead of assuming random initial orbits for satellites, they track the positions of their associated subhaloes. For cases where the subhalo containing a satellite can no longer be resolved following mass stripping, the position and velocity of the subhalo when it was last identified are used to compute a merger time-scale. This time-scale is then used in the same way as in the default scheme described above. The merger time-scale calculation assumes an NFW (Navarro, Frenk & White 1996) halo mass distribution to compute the orbital parameters of the satellite, combined with the analytical time-scale result of Lacey & Cole (1993). If a halo formation event occurs at a time after the subhalo is lost, a new merger time-scale for the satellite is calculated in the same way, using instead the position and velocity of the particle which was the most bound particle of the subhalo when it was last identified.

In this new GALFORM merger scheme, a satellite galaxy is not allowed to merge while it remains associated with a resolved subhalo. This treatment is similar to the scheme employed in the LGALAXIES model. The choice of merger scheme has an impact on the small-scale clustering, and contributed to the differences between the predictions of LGALAXIES and GALFORM reported by Contreras et al. (2013). The differences between the clustering predictions using the two types of merger scheme can be explained in terms of the concentration of the radial distribution of satellites, and also changes in the number of satellites. Requiring satellites to track their resolved
subhaloes, rather than computing an analytical merger time-scale as soon as a galaxy becomes a satellite, results in a more radially extended distribution of satellites (as demonstrated by Contreras et al.).

No model parameters have been recalibrated when using the new merger scheme. This would be likely to involve changing the strength of AGN feedback, and the time-scale for gas return from supernovae. We leave such calculation for a future paper. When comparing the model predictions using the new merger scheme to observational estimates of galaxy clustering as a function of stellar mass (in Section 5.2.3), we require that the model stellar mass functions reproduce those inferred from observations, through abundance matching (see Section 4.1.2). It is possible that making use of the subhalo mass at the time it was last identified results in shorter merger time-scales than would be computed if any subsequent mass stripping of the subhalo could be taken into account.

3 SED FITTING AND OBSERVATIONAL DATA

In theoretical models such as GALFORM, the galactic stellar masses are predicted from the SFHs. However, for observed galaxies, the stellar masses are not known directly but must be derived from observables. SED fitting is a popular technique for computing stellar masses. The amplitude of the 4000 Å break, \( r' \), and \( i' \) bands to identify candidates for spectroscopic follow-up (York et al. 2000). Li et al. (2006) estimated the clustering of galaxies as a function of stellar mass using the New York University Value-Added Galaxy Catalog (Blanton et al. 2005), which is based on the second SDSS data release (Abazajian et al. 2004). This catalogue has photometry covering 3514 deg\(^2\), and spectroscopy covering 2627 deg\(^2\) (about 85 per cent complete), for redshifts \( z \lesssim 0.3 \). Li et al. define their sample of galaxies with the magnitude and redshift limits given in Table 1, yielding a total of 196 238 galaxies. Subsets of this sample, defined in stellar mass, are used to study galaxy clustering.

3.2 Sloan Digital Sky Survey

The Sloan Digital Sky Survey (SDSS) uses photometry in the \( u, g, r, i, \) and \( z \) bands to identify candidates for spectroscopic follow-up (York et al. 2000). Li et al. (2006) estimated the clustering of galaxies as a function of stellar mass using the New York University Value-Added Galaxy Catalog (Blanton et al. 2005), which is based on the second SDSS data release (Abazajian et al. 2004). This catalogue has photometry covering 3514 deg\(^2\), and spectroscopy covering 2627 deg\(^2\) (about 85 per cent complete), for redshifts \( z \lesssim 0.3 \). Li et al. define their sample of galaxies with the magnitude and redshift limits given in Table 1, yielding a total of 196 238 galaxies. Subsets of this sample, defined in stellar mass, are used to study galaxy clustering.

Li et al. use the approach of Kauffmann et al. (2003) to estimate stellar masses. The amplitude of the 4000 Å break, \( D_{2600} \) (Balogh et al. 1999), and the strength of the Balmer H\( \alpha \) absorption line (Worthey & Ottaviani 1997) are measured from a spectrum obtained with a 3 arcsec diameter fibre. These measurements encode information about the age of the galaxy’s stellar population (\( D_{2600} \)), and can be used as indicators of whether recent star formation has

<table>
<thead>
<tr>
<th>Sample property</th>
<th>SDSS</th>
<th>GAMA</th>
<th>VIPERS</th>
</tr>
</thead>
<tbody>
<tr>
<td>Source</td>
<td>Li et al. (2006)</td>
<td>Farrow et al. (in preparation)</td>
<td>Marulli et al. (2013)</td>
</tr>
<tr>
<td>Galaxies used</td>
<td>157 887</td>
<td>50 770</td>
<td>17 100</td>
</tr>
<tr>
<td>Apparent magnitude limits</td>
<td>( 14.5 &lt; r &lt; 17.77 )</td>
<td>( r &gt; 19.8 )</td>
<td>( i &lt; 22.5 )</td>
</tr>
<tr>
<td>Median redshift</td>
<td>( \sim 0.1 )</td>
<td>0.19, 0.20</td>
<td>0.62, 0.62, 0.62</td>
</tr>
<tr>
<td>Calzetti redshift for comparison</td>
<td>0.089</td>
<td>0.17</td>
<td>0.62</td>
</tr>
<tr>
<td>( \alpha )</td>
<td>(-1.073 \pm 0.003)</td>
<td>(-0.35 \pm 0.18, -1.47 \pm 0.05)</td>
<td>(-0.95 \pm 0.03)</td>
</tr>
<tr>
<td>( 10 \log(M_\odot h^{-2}[Mpc^3]) )</td>
<td>(10 + \log_{10}(4.11 \pm 0.02))</td>
<td>(10.66 \pm 0.05 + \log_{10}(0.7^2))</td>
<td>(10.87 \pm 0.02 + \log_{10}(0.7^2))</td>
</tr>
<tr>
<td>( \Phi (10^{-7} h^3 \text{Mpc}^{-3} \text{dex}^{-1}) )</td>
<td>((20.4 \pm 0.1) / \ln(10))</td>
<td>((3.96 \pm 0.34, 0.79 \pm 0.23) / 0.7)</td>
<td>(1.42 \times 10^{0.8 \pm 0.03} / 0.7)</td>
</tr>
</tbody>
</table>
been predominantly quiescent, or due to bursts (Hδ). The stellar mass-to-light ratio in the $z$ band is estimated for each galaxy, by fitting $D_{4000}$ and Hδ to a Monte Carlo library of stellar populations, based on the Bruzual & Charlot (2003) SPS model. A Kroupa (2001) IMF is assumed in building the library, i.e. $x = -0.3$ for $M_* < 0.08 M_\odot$, $x = 0.3$ for $0.08 M_\odot < M_* < 0.5 M_\odot$, and $x = 1.3$ for $M_* > 0.5 M_\odot$ (see equation 1). Exponentially declining SFHs are used, with additional random bursts. The dust extinction applied in the best-fitting model is assumed for each galaxy, making use of a power-law attenuation curve, corresponding to a foreground dust screen (Charlot & Fall 2000). The stellar mass is then found by using the derived mass-to-light ratio in the $z$ band, combined with full $z$-band photometry (i.e. not limited by the fibre diameter). In this way, the mass-to-light ratio and dust attenuation derived within 1.5 arcsec of the galactic centre are extrapolated over the full galaxy. The masses estimated following Kauffmann et al. using spectral features have been shown to have a scatter of about 0.1 dex with respect to those obtained using SED fitting to broad-band photometry (Blanton & Roweis 2007; Li & White 2009).

The clustering results presented by Li et al. use six bins spaced logarithmically in stellar mass, covering $10^{8.69} < M_* < 10^{11.69}$. All but the highest mass bin correspond to samples which are volume limited in stellar mass, while each volume limited stellar mass bin covers a different redshift interval. However, the highest mass bin is flux limited, and thus incomplete in stellar mass. A correction is made by Li et al. to the correlation function computed for this bin, by weighting the contribution from each galaxy pair by the maximum volume over which they could be detected in the survey volume. They find that applying the same approach to flux limited samples for the lower stellar masses produces good agreement with the clustering results for the volume limited samples. Thus we assume that the results do not suffer from incompleteness in stellar mass due to magnitude limiting.

Li et al. find increasing clustering amplitude as a function of stellar mass, with a sharp increase above the characteristic stellar mass ($\sim 10^{10.6} h^{-2} M_\odot$), which is similar to the clustering trend they find when selecting galaxies by their $r$-band luminosity.

### 3.3 Galaxy and Mass Assembly Survey

The Galaxy and Mass Assembly Survey (GAMA) is a multiwavelength survey (far-UV to far-IR), with complete spectroscopy to $r = 19.8$ (Driver et al. 2011; Liske et al., in preparation; see Table 1). There are two clustering studies in GAMA focusing on stellar mass (Farrow et al., in preparation; Palamara et al., in preparation). For the present analysis, we compare to the clustering of galaxies in differential stellar mass bins as computed by Farrow et al. using the GAMA-II data. This is selected from the seventh SDSS data release (Abazajian et al. 2009), with sky coverage of 180 deg$^2$, out to $z < 0.5$, and spectroscopic completeness of >98 per cent.

To estimate stellar masses, Farrow et al. use an empirical relation based on the observed $g - i$ colour and absolute magnitude in the $i$ band, as fitted by Taylor et al. (2011) to galaxies from the second GAMA data release, with $r < 19.4$ (Liske et al., in preparation). To derive this relation, Taylor et al. implement broad-band SED fitting using the Bruzual & Charlot (2003) SPS model, with a Chabrier (2003) IMF, and exponentially declining SFHs. Extinction by dust is modelled using the Calzetti et al. (2000) law. The photometric bands used in the fitting are the GAMA $u$, $g$, $r$, $i$, and $z$. Taylor et al. use likelihood-weighting of all template SEDs, which can suppress discreteness effects due to the lack of interpolation between the small number of metallicities typically available in SPS models, improving on the common practice of taking the mode of the likelihood distribution (e.g. Mitchell et al. 2013).

Farrow et al. use samples selected in redshift and stellar mass to compute the correlation function. We compare to their intermediate redshift range (see Table 1), considering the stellar mass range $10^{9.3} < M_* h^{-2} M_\odot < 10^{11.5}$. They find that clustering amplitude increases with stellar mass. A decrease in clustering strength with redshift is noted for masses below $10^{10.5} h^{-2} M_\odot$, with no significant redshift evolution above this.

### 3.4 VIMOS Public Extragalactic Redshift Survey

The VIMOS Public Extragalactic Redshift Survey (VIPERS) consists of spectroscopic observations of galaxies selected using Canada–France–Hawaii Telescope Legacy Survey (CFHTLS) photometry (Guzzo et al. 2014). Marulli et al. (2013) present galaxy clustering as a function of stellar mass in the first VIPERS data release (Garilli et al. 2014). This data set has sky coverage of about 11 deg$^2$, with spectroscopic completeness of roughly 40 per cent. The redshift range sampled is $0.5 < z < 1.2$, and the magnitude limit is as given in Table 1. A selection in colour is used to exclude galaxies with $z < 0.5$, which is not completely efficient at selecting galaxies towards the lower limit of the surveyed redshift range, i.e. the selection does not correspond exactly to a step function at $z = 0.5$ (Garilli et al. 2014). However, in practice this has little impact on the number of galaxies recovered across the full redshift range, and is not important for our comparisons, according to tests with the model galaxies (see Section 4).

Marulli et al. use broad-band SED fitting to estimate the stellar masses of their sample of VIPERS galaxies, as described by Davidzon et al. (2013). The SPS model of Bruzual & Charlot (2003) is used, with a Chabrier (2003) IMF, i.e. the same choices as made by Taylor et al. (2011) for GAMA (see Section 3.3). Both the Calzetti et al. (2000) and Prevot–Bouchet (Prevot et al. 1984; Bouchet et al. 1985) dust attenuation laws are used, with the best-fitting option being chosen for each galaxy. Both exponentially declining and constant SFHs are used in the fitting: The photometric bands used are the following: the CFHTLS $u$, $g$, $r$, $i$, and $z$; the Wide-Field Infra-Red Camera (WIRCAM; Puget et al. 2004) $K$; the Galaxy Evolution Explorer (GALEX; Gil de Paz et al. 2007) far-UV and near-UV; the UKIRT Infra-Red Telescope Deep Sky Survey (UKIDSS; Lawrence et al. 2007) $J$, $H$, and $K$; and the Spitzer Wide-Area Infra-Red Extragalactic Survey (SWIRE; Lonsdale et al. 2003) 3.6 and 4.5 μm.

Samples selected in redshift and stellar mass are used by Marulli et al. to compute the correlation function. The clustering strength increases with stellar mass in each redshift range. We compare the model clustering predictions to the lower redshift interval.
considered by Marulli et al. (see Table 1), for stellar masses $M_\ast > 10^{9.5} \, h^{-2} \, M_\odot$.

4 ANALYSIS

Our analysis consists of two distinct components: (i) computation of the real space clustering as a function of stellar mass in the Gon14 and Lac14 GALFORM models at redshifts of 0.089, 0.32, and 0.62, and (ii) comparison of the projected model clustering as a function of stellar mass to observational data at redshifts of 0.089, 0.17, and 0.62 (i.e. the MS-W7 snapshots closest to the median redshifts of the observed galaxy samples; see Table 1). In (i), the lower and upper redshifts are the same as those in (ii), for comparing to SDSS and VIPERS data, respectively; however, the intermediate redshift (0.32) is chosen to be roughly evenly spaced in lookback time between these redshifts, covering 1.2–5.9 Gyr, rather than using the redshift of the comparison to GAMA (0.17). In (ii), we carry out SED fitting to obtain stellar mass estimates for the model galaxies, to allow a more robust comparison to the observations. The predictions obtained using the new subhalo dynamical friction (SDF) merger scheme are considered in both parts of our analysis.

For each model galaxy, GALFORM outputs the true stellar mass, real-space coordinates, peculiar velocity, and photometry including dust attenuation (see Section 2.1). Redshift space coordinates are computed by taking the line of sight as the third Cartesian axis and taking into account the peculiar motions along this axis (this is the distant observer approximation for projected clustering).

We convert the stellar masses predicted by GALFORM from units of $h^{-1} \, M_\odot$ to $h^{-2} \, M_\odot$ using $h = 0.704$ (see Section 2.2), to be consistent with the mass units of the observational data.

Apparent magnitude limits were imposed to match those of each survey, at certain stages in our analysis which will be indicated in Sections 4.1.2 and 4.2.2. These made use of SDSS $r$-band filter wavelength response data to match GAMA, and CFHTLS $i$-band filter response data to match VIPERS (see Table 1). In matching SDSS, we did not impose (apparent or absolute; cf. Table 1) magnitude limits, as the clustering results of Li et al. (2006) correspond to galaxy samples which are volume limited (see Section 3.2).

VIPERS uses a selection in colour to exclude galaxies with $z < 0.5$, as noted in Section 3.4, which leads to incomplete sampling of galaxies close to $z = 0.5$. Guzzo et al. (2014) note that the erroneous exclusion of galaxies ceases for $z \geq 0.6$. We have verified that this selection is indeed unimportant for the models by $z = 0.62$.

This section describes our adjustments to the model galaxy stellar masses (Section 4.1), and the computation of the correlation functions used to describe their clustering (Section 4.2).

4.1 Stellar masses

When computing clustering predictions to compare to observations, we considered three sets of model stellar masses: (i) the true stellar masses as predicted by GALFORM; (ii) estimates of the masses from SED fitting to the model photometry; and (iii) masses resulting from abundance matching to the stellar mass functions reported for the observed galaxies (as an adjustment following the SED fitting). In particular, we present the clustering results of the models with the standard merger scheme using (i), (ii), and (iii), and the results of the new SDF merger scheme using (iii). The determination of (ii) and (iii) will now be described.

4.1.1 SED fitting

Mitchell et al. (2013) implemented SED fitting (see Section 3.1) to estimate stellar masses for model galaxies, using broad-band photometry. These stellar mass estimates, when compared to the true values calculated in the model, can be used to investigate the influences of the various assumptions which are required in SED fitting (e.g. differences in the choice of IMF, the SFH, recycling of stellar mass back into the interstellar medium, the metallicities available in the SPS models, and dust attenuation), on the derived properties of the galactic population.

Following Mitchell et al., we carried out SED fitting for both GALFORM models, at each redshift, in order to obtain estimates of the stellar masses of the model galaxies. As noted in Section 3, the GAMA and VIPERS stellar masses are themselves derived through broad-band SED fitting (the GAMA stellar masses use an empirical formula based on this; Taylor et al. 2011), so our intention here is to carry out equivalent fitting procedures for the model galaxies. The SDSS stellar masses were derived by fitting to particular spectral features, rather than to broad-band photometry (see Section 3.2); we nonetheless carried out broad-band SED fitting for this comparison.

In all cases the Bruzual & Charlot (2003) SPS model was assumed, with a Chabrier (2003) IMF, exponentially declining SFHs, and the Calzetti et al. (2000) dust extinction law. The metallicities used to compute the template SEDs were matched to those used in the GAMA and VIPERS SED fitting by Taylor et al. (2011) and Davidzon et al. (2013), respectively, at the relevant redshifts. Taylor et al. use the full native metallicity grid of the Bruzual & Charlot SPS model, i.e. $Z \in \{0.0001, 0.0004, 0.004, 0.008, 0.02, 0.05\}$, while Davidzon et al. use $Z \in \{0.004, 0.02\}$. To compare to SDSS, the full set of metallicity values was employed in the fitting, as in the comparison to GAMA. Each template SED had a non-evolving metallicity, and we did not interpolate between the available SPS model metallicities, in keeping with the typical observational SED fitting procedures. All the photometric bands listed in Sections 3.3 and 3.4 were used in the SED fitting to match GAMA and VIPERS, respectively, with the exception of the UKIDSS K band, which was only included by Davidzon et al. in the absence of WIRCAM K-band magnitudes for VIPERS. The standard SDSS $u, g, r, i$ and $z$ filter set was used in the fitting to match SDSS. Taylor et al. compute likelihood-weighted stellar masses (see Section 3.3). We conformed to this for the fitting to match GAMA, but used the more standard approach of selecting the SED at the mode of the likelihood distribution for the fitting to match SDSS and VIPERS.

4.1.2 Schechter function matching

The logarithmic stellar mass function of galaxies $\Phi(M_\ast)$ is defined such that the number of galaxies per unit volume with stellar mass in the range $\log_{10}(M_\ast/M_{\text{ref}})$ to $\log_{10}(M_\ast/M_{\text{ref}}) + \, \sigma \, \log_{10}(M_\ast/M_{\text{ref}})$ is $dn = \Phi(M_\ast) \, d\log_{10}(M_\ast/M_{\text{ref}})$, for some mass unit $M_{\text{ref}}$. This is conventionally described using a Schechter (1976) function, where $\alpha$ is a characteristic mass, $\alpha$ is a power-law slope, and $\Phi$ is a normalization:

$$dn = \ln(10) \Phi \left( \frac{M_\ast}{M_*} \right)^{\alpha+1} \exp \left( -\frac{M_\ast}{M_*} \right) \, d\log_{10} \left( \frac{M_\ast}{M_{\text{ref}}} \right).$$

Differences between the numbers of model and observed galaxies in a given mass interval may give rise to discrepancies in the clustering results for samples selected by stellar mass, even if the underlying clustering signal is identical. Such mass function differences are dependent on the details of the model physics, combined
with the procedure for estimating stellar masses through SED fitting. In order to eliminate any differences between the model and observationally inferred stellar mass functions, the Schechter functions representing the stellar mass functions of the galaxies in each observational sample (see Table 1) were imposed on the model galaxies at the corresponding redshifts. This process is equivalent to rescaling the estimated stellar mass of each GALFORM galaxy, in order to make the model mass functions match the observational results. The procedure used was to match the shape and normalization of the target Schechter function, while maintaining the ordering of the model galaxies in estimated stellar mass from SED fitting, as follows. Given the lowest stellar mass of interest (for the clustering samples, see Section 3), and the Schechter function fitted to the measurement for the observed galaxies, we generated the expected number of galaxies in the simulation volume by randomly sampling stellar masses consistent with this mass function. The generated masses were assigned in order to the GALFORM galaxies, such that the highest generated mass was ascribed to the GALFORM galaxy with the highest mass estimate from SED fitting, and so on. It is important to note that the galaxy formation models considered here have not been calibrated to reproduce observationally inferred stellar mass function data. They have, however, been calibrated to match the local K-band luminosity function.

Measuring $\Phi(M)$ requires one to know the number of galaxies within some stellar mass range of interest, in some known volume. Difficulties in achieving this arise from the fact that galaxy surveys are defined by apparent magnitude limits, i.e. a sample of survey galaxies defined by stellar mass limits will inherently also be restricted to some range in apparent magnitude. It is, however, still possible to identify samples of galaxies which are complete, i.e. volume limited, in stellar mass (e.g. the clustering samples of Li et al. 2006). The level of incompleteness in stellar mass, i.e. the fraction of missing galaxies, in a sample defined by stellar mass, depends on the apparent magnitude limits of the survey and the redshift range of interest.

In the typical case, where we are concerned with a faint apparent magnitude limit, it is useful to be able to estimate a lower mass threshold above which the measured mass function can be considered to be ‘reliable’ (i.e. the same as what would have been measured with no magnitude limit, for a volume limited sample). Pozzetti et al. (2010) describe the method used by Davidzon et al. (2013) to estimate such a threshold mass for VIPERS. At our comparison redshift (see Table 1), this VIPERS threshold value is approximately $10^{9.6} h^{-2} M_{\odot}$, which corresponds to the mass below which the faint limit causes more than about 20 per cent of galaxies to be missed, in a given mass interval. The VIPERS Schechter function specified in Table 1 has been fitted by Davidzon et al. only for masses above the reliability threshold defined in this way. Clearly the i-band faint limit has an important influence on the completeness of the stellar mass function at this redshift, for the lowest masses of interest (note that the minimum mass we consider for the clustering calculations using these galaxies is $10^{10.5} h^{-2} M_{\odot}$). In light of this, we imposed the i-band faint limit on the model galaxies before matching the VIPERS mass function at this redshift, in order to reproduce the ‘underestimated’ mass function measurement at lower masses, i.e. the incomplete mass function of the galaxies actually used in the clustering analysis of Marulli et al. (2013).

The double Schechter function representing the GAMA stellar mass function has been fitted for $z < 0.06$ (Baldry et al. 2012), whereas the clustering results of Farrow et al. (in preparation) considered here use galaxies with $0.14 < z < 0.24$ (see Table 1). Baldry et al. use the SED fitting procedure of Taylor et al. (2011) to estimate stellar masses, supplemented by the corresponding empirical relation of Taylor et al. for a small number of galaxies with missing photometry (see Section 3.3). Baldry et al. compare their stellar mass function to that obtained by Pozzetti et al. (2010) using the Cosmic Evolution Survey (zCOSMOS; Lilly et al. 2007) for $0.1 < z < 0.35$, using similar SED fitting for the stellar mass estimation, and find good agreement between the two measurements. Furthermore, we have verified that the stellar mass function does not evolve significantly when using the sample of galaxies considered by Farrow et al. (0.14 < z < 0.24), with respect to the Baldry et al. measurement. The Baldry et al. mass function fit can be regarded as complete for the masses of interest here, thanks to being constrained at low redshift (the fitting considered the mass function as measured for $M_* > 10^{10.2} h^{-2} M_{\odot}$, while the clustering data we compare to are for $M_* > 10^{11.2} h^{-2} M_{\odot}$). Thus, we matched the mass functions of the GALFORM models to the double Schechter fit, while keeping all the model galaxies, and only afterwards imposed the r-band faint limit on the models (see Table 1). In this way, the complete model stellar mass functions were made to reproduce the complete observationally inferred mass function, before introducing the relative incompleteness due to the particular selection relevant for comparison to the clustering data of Farrow et al.

The SDSS mass function Schechter fit was derived directly from the second SDSS data release (Abazajian et al. 2004; Wang et al. 2006), using r-band apparent magnitude and redshift limits the same as for the clustering data of Li et al. (2006), but without the additional absolute magnitude limits imposed for the clustering samples (see Table 1). We assumed the Wang et al. fit to be approximately complete in stellar mass, and thus imposed it upon the model galaxies without applying any magnitude limiting, as required for comparison to the volume limited clustering samples of Li et al. (see Section 3.2). It should be noted, however, that the measured mass function is likely to be significantly suppressed by the SDSS faint r-band limit for $M_* < 10^{12.2} h^{-2} M_{\odot}$ (Baldry, Glazebrook & Driver 2008; Li & White 2009). This incompleteness has minor implications for our comparison to the Wang et al. mass function data and Schechter fit (for $M_* > 10^{10.6} h^{-2} M_{\odot}$) in Section 5.2, but is not important for our comparison to the volume limited clustering data; except where the mass function is imposed, in which case any errors introduced by the assumption of completeness in the Schechter fit should only have an impact on the lowest masses considered.

In Sections 5.2.2 and 5.2.3 we make use of the mass function matching to compare the results of the two GALFORM merger schemes to the observational clustering data (as outlined at the beginning of Section 4.1). In this way, we force the models using either scheme to reproduce the same mass function, and thus the differences in the clustering results using the two schemes, when compared in this way, are not due to any differences in the predicted stellar mass function which are introduced by changing to the new merger scheme.

4.2 Clustering

We now define the correlation function $\xi(r)$, and the projected correlation function $w_p(r)$, and give details of their computation.

4.2.1 Correlation function

For a cosmologically representative volume $V$, the probability $dP$ of finding a galaxy in a volume element $dV$, at a comoving distance
from another galaxy in a volume element $dV_2$, defines the spatial
two-point autocorrelation function $\xi(r)$ such that
\[
d^p = n_V^2 \left[ 1 + \xi(r_{12}) \right] dV_1 \, dV_2, \tag{3}
\]
where $n_V$ is the mean number density of galaxies within $V$ (Peebles 1980).
Thus following Rivolo (1986),
\[
1 + \xi(r) = \frac{1}{n_V^2} \frac{N(r)}{V(r)}, \tag{4}
\]
where $N(r)$ is the number of pairs with separation $r$, computed by
summing over the number of pairs involving each galaxy in the volume
individually (so this is twice the number of independent pairs),
considering a spherical shell of radius $r$ and volume $V(r)$. A
simple appreciation of the uncertainty in $\xi(r)$ can be gained from
considering Poisson statistics (Iovino & Shaver 1988; Martinez et al.
1993). The error on $\xi(r)$ is then $\Delta \xi(r)$ such that
\[
\Delta \xi(r) = \left( \frac{2}{n_V^2} \frac{1 + \xi(r)}{V(r)} \right)^{1/2}. \tag{5}
\]
It is common to fit a power law to $\xi(r)$, parametrized by a corre-
lation length $r_0$ and slope $\gamma$ (Peebles 1980):
\[
\xi(r) = \left( \frac{r}{r_0} \right)^{-\gamma}. \tag{6}
\]
Here $r_0$ characterizes the clustering ‘strength’, where $\xi(r_0) = 1$. This
parametrization is not suitable for describing the precise clustering
measurements which are possible today over a wide range of scales,
but we can make use of it to describe the galaxy clustering over small
ranges of pair separations.

$\xi(r)$ was computed for the models at redshifts $z \in \{0.089, 0.32,$
$0.62\}$ (using the same stellar masses, with no magnitude limits),
with and without the new merger scheme, in three equally spaced bins
in stellar mass. Pair separations in the range $0.1 < r [h^{-1}$ Mpc $] < 30$
(the choice of the large-scale limit will be discussed in Section 4.2.3)
were divided into 30 bins of equal logarithmic width. $\xi(r)$ was
computed using the pair counts binned in $r$, following equation (4),
where $V = L_{\text{box}}^3$ (see Section 2.2).

The power law given in equation (6) was fitted to each $\xi(r)$, using
the $\Delta \xi(r)$ from equation (5) to weight the fit. We preferred to fit over
a relatively small range in $r$, where $\xi(r)$ is close to being an exact
power law, in the neighbourhood of $\xi(r) = 1$, such that $r_0$ relates
closely to the ‘true’ correlation length. In this way, bins within
$3 < r [h^{-1}$ Mpc $] < 10$ were used to fit the power law for each $\xi(r)$.
This range in $r$ samples the two-halo term in the correlation function,
that is, the separations considered relate to galaxies in different
haloes. As such, the fitted power laws are insensitive to clustering on small
scales (one-halo term), and thus we consider them using only the
standard merger scheme.

Our results for $\xi(r)$ using the two GALFORM models, comparing
the two merger schemes, are presented in Section 5.1.

4.2.2 Projected correlation function

The distance to a galaxy moving with the cosmic expansion can be
inferred through measuring its recession velocity (redshift) and
using Hubble’s law. If the galaxy has some peculiar velocity relative
to the Hubble flow, this will perturb the determined distance. As
such, it is necessary to distinguish between real space and redshift
space coordinates, i.e. true spatial positions, and those determined
through measuring recession velocities, respectively (Kaiser 1987).

For observed galaxies, the redshift space correlation function may
be computed. This encodes information about the peculiar motions
of galaxies on different separation scales. Through considering
the components of the pair separation orthogonal and parallel to the line
of sight ($\sigma$ and $\pi$, respectively), we can define the two-dimensional
correlation function $\xi(\sigma, \pi)$ by analogy with equation (4). We
may integrate $\xi(\sigma, \pi)$ along the line of sight, defining the projected
correlation function $w_p(\sigma)$ (Davis & Peebles 1983),
\[
w_p(\sigma) = \int_{-\sigma_{\text{max}}}^{\sigma_{\text{max}}} \xi(\sigma, \pi) \, d\pi. \tag{7}
\]
This statistic is traditionally used to describe the real space clustering
of galaxies drawn from observational surveys, as redshift space distortions
do not influence $w_p(\sigma)$. This is true provided that $\sigma_{\text{max}}$ is sufficiently
large, such that the integral includes all correlated galaxy pairs,
including the redshift-distorted coordinates (Davis & Peebles 1983; Norberg et al. 2009).

Norberg et al. show using simulations that $w_p(\sigma)$ is sensitive to the choice
of $\sigma_{\text{max}}$, with differences between the projected clustering
on large scales for $\sigma_{\text{max}} [h^{-1}$ Mpc $] \in \{30, 64\}$.

As the correlation function, $w_p(\sigma)$, is the corresponding
power law for $w_p(\sigma)$ is given by (Davis & Peebles 1983)
\[
w_p(\sigma) = \sigma^{1-\gamma} \frac{\Gamma(1/2) \Gamma(\gamma - 1/2)}{\Gamma(\gamma/2)} \sigma_{\text{max}}, \tag{8}
\]

This result is used in this paper only for illustrative purposes.

$w_p(\sigma)$ was computed in real and redshift space for the two
GALFORM models at each redshift $z \in \{0.089, 0.17, 0.62\}$, for each
of the four variants identified in Section 4.1, i.e. the default models
with their true and estimated masses, and the models with masses
resulting from matching the observational mass functions (using
both the new and default merger schemes). At each redshift the
appropriate magnitude limits were imposed (see the beginning of
Section 4). The stellar mass binning was matched to that of each
survey. Separations orthogonal to the line of sight were computed with
the same binning as the $r$ values for $\xi(r)$ (see Section 4.2.1). Line
of sight separations were measured in bins of width $\Delta \pi = 1 h^{-1}$ Mpc,
using the distant observer approximation (as noted at the beginning
of Section 4). $\xi(\sigma, \pi)$ was evaluated for each bin of the $\sigma$ versus
$\pi$ grid, using equation (4), where the volume $V_i$, considered was
instead that of a cylindrical shell of radii set by the $\sigma$ bin edges
and length $\Delta \pi$. Discretizing equation (7), $w_p(\sigma) = \sum_{i \in \sigma} \xi(\sigma_i, \pi_i) \Delta \pi$
was computed for bin centres $\pi_i$, such that $|\pi| < \pi_{\text{max}} [30 h^{-1}$ Mpc,
see Section 4.2.3.

The results of our comparison of the model clustering to the
observational $w_p(\sigma)$ data are presented in Section 5.2.
To establish the pair separations where the box size becomes important, we computed the projected correlation function along three mutually orthogonal lines of sight. This was carried out for each redshift and stellar mass range of interest in this paper, in both real and redshift space.

We found the scatter in \( w_{p}(\sigma) \) for \( \sigma < 30 \, h^{-1} \) Mpc, around the mean for the three projections, to be about 0.01 dex. For \( \sigma > 30 \, h^{-1} \) Mpc, the scatter rose sharply with increasing \( \sigma \), reaching roughly 0.1 dex by \( \sigma \sim 50 \, h^{-1} \) Mpc. These results for the \( w_{p}(\sigma) \) scatter across projections are largely insensitive to the choice of line of sight integration limit, over the range \( 30 < \pi_{\text{max}} [h^{-1} \) Mpc] < 50.

Based on this test, we used an upper limit of \( 30 \, h^{-1} \) Mpc for \( \sigma \) when computing \( w_{p}(\sigma) \), and for \( r \) when computing \( \xi(r) \). \( \pi_{\text{max}} = 30 \, h^{-1} \) Mpc was used when comparing to each survey. We have checked that using the particular \( \pi_{\text{max}} \) values given in Table 1 for each survey instead (which are larger for the SDSS and GAMA measurements) does not have a significant impact on our comparisons. Our upper limit to pair separations is the same as that noted by Orsi et al. (2008) as the largest separation where the galactic and dark matter \( \xi(r) \) are related by a scale-independent bias in the Millennium Simulation volume. Gonzalez-Perez et al. (2011) also found this scale to be the upper limit for describing the redshift space correlation function boost with the Kaiser (1987) formalism, due to the finite size of the simulation box.

5 RESULTS AND DISCUSSION

Our results are now presented for the real space \( \xi(r) \) computed using the \texttt{GALFORM} models as a function of true stellar mass and redshift (Section 5.1), and the comparisons to the \( w_{p}(\sigma) \) observational clustering estimates (Section 5.2).

5.1 Predicted real space clustering

We first consider the results of the default models (Section 5.1.1), and then examine the influence of the choice of satellite merger scheme on the small-scale clustering (Section 5.1.2).

5.1.1 Standard models

The galactic stellar mass function \( \Phi(M_{*}) \) as a function of redshift is shown in Fig. 1, as predicted using the Gon14 and Lac14 \texttt{GALFORM} models. Fig. 2 shows the real space \( \xi(r) \) computed as a function of stellar mass and redshift. Fig. 3 shows the parameters \( r_{0} \) and \( \gamma \) of the power laws fitted to each \( \xi(r) \) bin.

As shown in Fig. 1, the evolution of the shape of \( \Phi(M_{*}) \) as a function of redshift is fairly similar between the models. The number density above the knee of the mass function decreases with increasing redshift in each model. There are more high-mass galaxies in the Gon14 model than in the Lac14 model (this is similar to the differences between the predicted K-band luminosity functions). This regime roughly corresponds to the highest mass bin for the correlation functions shown in Fig. 2 (as indicated in Fig. 1).

For clarity, we divide the predicted correlation functions by a reference power law (see Fig. 2 caption). The model clustering predictions shown in Fig. 2 are close to a power law over only a limited range of scales. There is a clear transition between the one-halo term on small scales and the two-halo term on large scales (at \( r \sim 2 \, h^{-1} \) Mpc). The two-halo term has the same shape as the large-scale dark matter \( \xi(r) \), where the galaxy clustering bias with respect to the dark matter increases with both redshift and stellar mass. The clustering in the two models is indistinguishable on large scales (i.e. for pair counts between haloes), but there are differences between the model predictions on smaller scales (\( r \lesssim 1 \, h^{-1} \) Mpc).

There is a general trend of clustering amplitude increasing with stellar mass, on all separation scales, at each redshift. Slight increases in the amplitude of the clustering on large scales are seen with decreasing redshift, for a fixed mass bin. The power-law fits shown in Fig. 2 ignore clustering on small and very large scales. The fit parameters, as shown in Fig. 3, indicate increasing clustering strength \( r_{0} \) with both increasing stellar mass and decreasing redshift, where the differences between redshifts become less significant for higher stellar masses. The slope \( \gamma \) is fairly constant, but exhibits a weak increase with increasing stellar mass. The redshift evolution of the clustering on small scales in Fig. 2 (which is sensitive to pair separations within haloes) is particularly dramatic for low stellar masses.

5.1.2 New satellite merger scheme

In the new merger scheme, satellite galaxies track the positions of their associated subhaloes, only making use of an analytically computed merger time-scale once the subhalo can no longer be resolved in the simulation, as described in Section 2.4. Fig. 4 shows the stellar mass functions of the Gon14 and Lac14 models at redshift \( z = 0.089 \), with and without the new merger scheme, decomposed into the contributions from central and satellite galaxies. The same trends are seen at higher redshifts. Changes in the total \( \Phi(M_{*}) \) due to the new merger scheme are significant for the highest masses (top panel of Fig. 4), where for both models the amplitude of the high-mass end increases when switching to the new scheme. Note that the model parameters have not been recalibrated on adopting the new merger scheme. The rise in the galaxy abundance in the Lac14 model at these masses is greater than for the Gon14 model, such that with the new scheme, the two model mass functions become very similar for high masses. The increases in the numbers of the most massive galaxies are due to more mergers taking place in the new scheme, which thus transfer more mass to central galaxies, as can be seen in the middle panel of Fig. 4. There is a smaller rise at high masses in the satellite mass function for both models, which again is more significant for the Lac14 model (lower panel). The source of the growth of higher mass galaxies here is the earlier
Figure 2. Predicted correlation function of galaxies $\xi(r)$ as a function of redshift $z$ (columns) and stellar mass $M_*$ (rows), computed using the Gon14 and Lac14 GALFORM models. For clarity, each $\xi(r)$ has been divided by a reference power law $\xi_{\text{ref}}(r)$, with parameters $r_0 = 5 \, h^{-1} \text{Mpc}$ and $\gamma = 2$ (horizontal line; see equation 6). The dotted lines are power-law fits to each GALFORM $\xi(r)$, for $3 < r \, [h^{-1} \text{Mpc}] < 10$, with the parameters shown in Fig. 3. The black curves show $\xi(r)$ for the dark matter at each redshift, computed as the Fourier transform of the MS-W7 linear theory power spectrum.

assimilation of lower mass satellites. Mergers on to a galaxy cease when it becomes a satellite of a more massive galaxy.

The changes in $\xi(r)$ for the two GALFORM models when using the new merger scheme are displayed in Fig. 5, at redshifts of 0.089, 0.32, and 0.62. The lines show the results with the SDF merger scheme, divided by the predictions using the original merger scheme, for each model, stellar mass interval, and redshift. The clustering differences with respect to the standard merger scheme are larger for the lower redshifts. Generally there is little change in the clustering on large scales. There are significant decreases in the small-scale clustering amplitude ($r \lesssim 1 \, h^{-1} \text{Mpc}$) for all but the highest mass, highest redshift, data. The intermediate-mass bin shows the largest decreases in $\xi(r)$ on small scales, although these are similar to the reductions in amplitude for the lower mass bin. There is significantly less change in the highest mass bin, relative to that in the other mass bins, at a given redshift. Changes in both the radial distribution and number of satellites contribute to the changes in the small-scale clustering when using the new scheme (see lower panel of Fig. 4, and Section 2.4).

5.2 Comparisons to observational data

We now present the results of our comparisons to observational measurements of the clustering of galaxy samples selected by stellar mass. First we consider the changes induced in the model predictions through stellar mass estimation using SED fitting (Section 5.2.1), then discuss rescaling of the estimated model mass functions to reproduce observational mass function measurements (Section 5.2.2), and present the changes in the clustering predictions resulting from using the new SDF merger scheme in GALFORM (Section 5.2.3). In Section 5.2.4 we comment on the need for SED fitting in order to make a fair comparison to observational clustering measurements, even when carrying out mass function abundance matching. Finally, we discuss some uncertainties relevant to these comparisons to observational clustering results (Section 5.2.5).

5.2.1 True and estimated stellar masses

Mitchell et al. (2013) show that for some ranges in stellar mass, the process of SED fitting to the broad-band photometry of model galaxies, as predicted by GALFORM, introduces a roughly constant offset in estimated stellar mass with respect to true model masses, combined with associated scatter (see Section 4.1.1). However, for GALFORM galaxies with strong dust attenuation, SED fitting tends to significantly underestimate the stellar mass, leading to important systematic differences between the distributions of true and estimated masses, which cannot be described by a constant mean offset and scatter. This phenomenon is relevant for model galaxies with true stellar mass $\gtrsim 10^{10} \, h^{-2} \, M_\odot$.

It is important to note that the IMF assumed in a galaxy formation model influences mass-to-light ratios. The predicted luminosity functions are directly sensitive to the IMF, and are used to calibrate
the models, which in turn affects the stellar masses. Within the machinery of SED fitting, the choice of IMF corresponds to a roughly constant offset in the resulting stellar masses.

The Chabrier (2003) IMF used in our SED fitting matches the GAMA and VIPERS stellar masses, but not the Kroupa (2001) IMF of the SDSS data (see Sections 3 and 4.1.1). Therefore we multiplied the mass estimates by 1.1 following the SED fitting to match SDSS. This factor is that found by Davidzon et al. (2013) as the mean offset between the stellar masses derived using Chabrier and Kroupa IMFs in their SED fitting (see also Baldry et al. 2008).

The stellar mass functions of the Gon14 and Lac14 models are shown in Fig. 6, at redshifts of 0.089, 0.17, and 0.62, alongside the relevant observationally inferred mass function data from SDSS, GAMA, and VIPERS, respectively. The true and estimated mass functions are shown for each model using solid and dotted lines, respectively (the dashed lines show the estimated mass functions using the new merger scheme, which will be discussed in Section 5.2.3). For the comparison to SDSS (Fig. 6a), no magnitude limits have been imposed on the model galaxies, because the clustering measurements of Li et al. (2006) are for samples which are complete in stellar mass (see Section 3.2). When comparing to GAMA and VIPERS (Figs 6b and 6c), we imposed the apparent magnitude limits relevant for comparison to the clustering data of Farrow et al. (in preparation) and Marulli et al. (2013), respectively (see Table 1 and Section 4).

The Baldry et al. (2012) GAMA stellar mass function fit can be regarded as complete in stellar mass over the range of masses shown in Fig. 6(b), as it was constrained for \( z < 0.06 \) (see Section 4.1.2). Therefore we also show the estimated model stellar mass functions without applying the GAMA magnitude limit in Fig. 6(b), as these are appropriate to compare to the Baldry et al. data over the full range of masses shown (these lines are shown with the same formatting as the magnitude limited versions). Thus the differences due to the magnitude limit can be seen where the dotted lines of a given colour do not overlap (the lower, magnitude limited, dotted lines are close to the solid true model stellar mass lines at the lowest masses, following a reduction in amplitude of roughly 0.4 dex due to the magnitude limit). In this way, it can be seen that the faint \(-r\)-band limit of GAMA influences the completeness of the estimated model stellar mass functions for \( M_\star \lesssim 10^{10} h^{-2} M_\odot \) at \( z = 0.17 \). The vertical (dotted or dashed) black lines in Fig. 6(b) indicate a threshold stellar mass \( M_{\text{thresh}} \) for each model, above which the GAMA magnitude limit causes 1 per cent of the galaxies in the
The true model stellar mass functions can change significantly when carrying out SED fitting. An effect which can be seen at each redshift is that the knee of the estimated mass function is smoothed out and suppressed with respect to the true mass function, for both models. The changes at the knee become larger with increasing redshift, and can be traced to the influence of dust attenuation (see Mitchell et al. 2013). Additionally, scatter in the estimated masses at a given true mass will cause \( \Phi(M_t) \) to rise for masses above this, particularly near the high-mass end, due to the rapid changes in number density as a function of stellar mass. This Eddington bias conspires with the systematic effects of dust attenuation to dictate the shape of the estimated mass function.

The projected correlation functions computed using the model galaxies, for samples selected by stellar mass, are shown in Figs 7–9, alongside the relevant observational measurements, for the comparisons to SDSS, GAMA, and VIPERS, respectively. In each case, column (a) shows the model clustering predictions when using the true stellar masses, and column (b) shows the results when using the estimated masses from SED fitting. These masses correspond to the true and estimated mass functions shown in Fig. 6, where the relevant \( w_p(\sigma) \) mass bin edges are indicated by vertical lines. The results shown in columns (c) and (d) of Figs 7–9 will be discussed in Sections 5.2.2 and 5.2.3, respectively.

Switching from true to estimated model stellar masses can change the clustering predictions significantly, comparing columns (a) and (b) in Figs 7–9. More massive haloes are more strongly clustered in hierarchical cosmologies (e.g. Cole & Kaiser 1989), and the stellar mass of a galaxy is closely related to its host halo mass (e.g. Guo et al. 2010). Therefore the changes in the clustering predictions when carrying out SED fitting can be understood in terms of model galaxies in dark matter haloes of certain masses being transferred across the stellar mass bin boundaries. The Gon14 and Lac14 models predict relatively similar clustering in most cases when using true stellar masses, but the SED fitting tends to significantly increase the differences between the results for the two models. The changes in the clustering predictions of the Gon14 model when switching to estimated masses are often larger than for the Gon14 model. It is likely that this is due to a combination of different factors, such as the different SPS models assumed and levels of dust extinction calculated in the two models. The SED fitting tends to decrease the clustering amplitude in the Lac14 model, but tends to increase that predicted in the Gon14 model. In general, switching to estimated masses has a larger impact on the clustering on small scales (\( \sigma \lesssim 1 \) h\(^{-1}\) Mpc), such that the shape of \( w_p(\sigma) \) is modified.

Comparing to SDSS, the agreement with the clustering measurements of Li et al. (2006) is improved or similar when switching from true to estimated stellar masses, for both GALFORM models (see Figs 7a and 7b). For both models in most mass bins, the clustering amplitudes using true or estimated masses are higher over the range of separations shown than inferred from SDSS. However, there is somewhat better agreement for the highest masses, in particular, when using estimated masses in the second highest (for Lac14) and highest (for Gon14 and Lac14) mass bins.

In Fig. 8a, the clustering predictions using true stellar masses are similar for the two models, except that the Lac14 model predicts relatively low clustering amplitude in the lower mass bin. There is reasonable agreement with the large-scale GAMA clustering measurements of Farrow et al. (in preparation), especially in the higher mass bin. Both models predict stronger clustering on the smallest scales than is inferred observationally; this difference is most significant for the Gon14 model in the lower mass bin. The Lac14 model clustering predictions yield improved agreement with the GAMA measurements when switching to estimated masses. However, the corresponding increases in the clustering amplitudes of the Gon14 model lead to poorer agreement with the GAMA data than found with the true stellar masses (see Figs 8a and 8b).

Using true stellar masses, the model predictions agree with the VIPERS clustering measurements of Marulli et al. (2013) on intermediate scales (\( \sigma \sim 1 \) h\(^{-1}\) Mpc) in the two lower mass bins in Fig. 9a), where they tend to overpredict the clustering amplitude on smaller and larger scales. In the highest mass bin, both models agree reasonably well with the observational measurements on small and large scales, but predict lower clustering amplitude than observed on intermediate scales. Switching to the estimated masses (see Fig. 9b), improves the large-scale agreement with the VIPERS data for the Lac14 model, in the two lower mass bins. Otherwise, the level of agreement between the model predictions and observational measurements on large scales is similar or reduced when switching to the estimated masses. On small scales, switching to estimated masses worsens the agreement with the VIPERS clustering data for each model and mass bin, except for the Lac14 model in the intermediate bin, where the level of agreement is slightly increased. These changes result in the models with estimated masses yielding clustering predictions which encompass the VIPERS data on small scales, with the clustering amplitude being lower for the Lac14 model, and higher for the Gon14 model.
It is clear from the results shown in columns (a) and (b) of Figs 7–9 that carrying out SED fitting to the broad-band photometry of model galaxies to recover estimates of their stellar masses can have a significant impact on the model clustering predictions for samples selected by stellar mass. Thus it is important to implement this procedure to obtain a reliable comparison between the clustering of galaxies as a function of stellar mass, as measured for observed galaxies and predicted by theoretical models.

5.2.2 Abundance matching of stellar mass functions

Column (c) of Figs 7–9 shows the projected correlation functions resulting from carrying out abundance matching of the estimated model stellar mass functions to Schechter (1976) function fits to the mass function data measured using each survey (SDSS, GAMA, and VIPERS, respectively). The mass function fits are shown alongside the model predictions in Fig. 6 (see Section 4.1.2 for details of the matching procedure). The differences between the model clustering predictions shown in columns (b) and (c) of Figs 7–9 are thus entirely due to the model mass estimates from SED fitting being rescaled to reproduce the observationally inferred mass function fits.

Looking at the comparison to the SDSS results of Li et al. (2006) in Figs 7(b) and 7(c), the mass function matching leads to a suppression of the model clustering amplitude for both models (in most mass bins). This leads to improved, or unchanged, agreement with the observational results; except for the Lac14 model in both the second highest mass bin (where poorer agreement is found due to a rise in clustering amplitude), and in the highest mass bin (where the model \( \Phi(M) \) becomes lower in amplitude than measured from SDSS). However, the changes in the clustering predictions due to imposing the SDSS mass function are typically relatively minor, and in general there remains excessive clustering on both small and large scales with respect the SDSS measurements in Fig. 7(c).

Comparing to the measurements of Farrow et al. (in preparation) using GAMA data (see Figs 8a and 8c), imposing the observationally inferred mass function causes the clustering amplitude to be suppressed for both models in the lower mass bin (where the most significant impact is on the Gon14 model). This leads to improved agreement with the GAMA clustering results for the Gon14 model, and similar agreement for the Lac14 model, in this mass bin. In the higher mass bin, imposing the mass function has very little influence on the Lac14 model clustering prediction, but slightly boosts the amplitude of the small-scale clustering in the Gon14 model, exacerbating the discrepancy with the GAMA measurement on small scales.

It can be seen in Figs 9(b) and 9(c) that the mass function matching leads to improved agreement with the VIPERS data in the lower mass interval for the Lac14 model, and worsened agreement for the Gon14 model, as the clustering amplitude rises in both models. The matching does not have a significant impact on the clustering predictions of either model in the intermediate-mass interval. Both models exhibit suppression in the clustering amplitude due to the mass function matching for the highest masses, improving the agreement with the VIPERS data for the Gon14 model, and reducing the level of agreement for the Lac14 model. Thus there persist notable differences between the model predictions and the VIPERS measurements following the abundance matching, in Fig. 9(c). The Gon14 model predicts excessive clustering with respect to the VIPERS results on both small and large scales in the two lower mass bins. The Lac14 model predicts lower clustering.
Galaxy clustering as a function of stellar mass

Figure 7. Projected correlation function of galaxies $w_p(\sigma)$ as a function of stellar mass $M_\star$ (as labelled in each row) at redshift $z = 0.089$, computed using the Gon14 and Lac14 GALFORM models, in real (first column only) and redshift space. The SDSS measurements of Li et al. (2006) are shown (black lines with error bars). Each $w_p(\sigma)$ has been divided by a reference power law $w_{\text{ref}}(\sigma)$, with parameters $r_0 = 5 \, h^{-1} \, \text{Mpc}$ and $\gamma = 2$ (horizontal line; see equation 8). Half the standard number of $\sigma$ bins are shown for the highest mass interval, due to the small number of galaxies. Columns (a) and (b) show the default models with the true and estimated masses, respectively. The remaining columns show the results of matching the SDSS stellar mass function (as an adjustment following the SED fitting), for (c) the default models, and (d) the models using the new SDF merger scheme.
5.2.3 New satellite merger scheme

Thus far, we have compared the model clustering predictions to observational results using the default treatment of satellite galaxies in GALFORM. In the default scheme, satellite galaxies are assumed to enter the main halo on random orbits (independently of the orbit of their associated subhalo), and merge with the central galaxy upon the elapse of an analytically determined dynamical friction merger time-scale. We now consider the results obtained when using a new merger scheme, in which satellite galaxies track their associated subhaloes until these are no longer resolved, and only then is an analytic merger time-scale computed (see Section 2.4).

Using the standard merger scheme (looking in particular at column (c) of Figs 7–9), the Gon14 model predicts higher small-scale clustering than seen in the observational data. The Lac14 model predicts higher small-scale clustering than measured from SDSS in Fig. 7(c), but is in agreement with (or predicts lower clustering amplitude than) the small-scale results from GAMA and VIPERS (Figs 8c and 9c, respectively). It is noteworthy that both models tend to predict higher clustering amplitudes on small scales than measured from SDSS.

Stronger small-scale clustering in the models, with respect to observational results, may be interpreted as an excess of satellite galaxies in the models, or too concentrated a radial distribution of satellites (Contreras et al. 2013). In order to reconcile the model predictions with observations, it is possible that a more detailed study of the disruption and mergers of satellite galaxies is needed (Henriques, Bertone & Thomas 2008; Kim et al. 2009). Another possibility is to make use of the substructure information available in dark matter simulations, i.e. to employ our new satellite merger scheme, which has an impact on small-scale clustering (see Sections 2.4 and 5.1.2).

The estimated stellar mass functions of the Gon14 and Lac14 models when using the new SDF merger scheme are shown as dashed lines in Fig. 6 (where they are labelled Gon14-sdf and Lac14-sdf). Comparing these lines to the estimated mass functions using the standard merger scheme (dotted lines), it can be seen that the most significant differences are at the highest masses. At each redshift, both models have more high-mass galaxies when using the new scheme, where the differences between the predictions for the two schemes for a given model emerge roughly at the knee of the mass function. For the true masses, the mass functions for both models using the new scheme are very similar to each other at the highest masses, for each survey comparison redshift (to the level shown in Fig. 4 at the SDSS comparison redshift); and yet the estimated mass functions shown in Fig. 6 are distinct at high masses for the Gon14 and Lac14 models using the new scheme.

Column (d) of Figs 7–9 shows the projected clustering predictions of the GALFORM models using the new merger scheme. These results use stellar masses estimated through SED fitting, followed by matching to the observational stellar mass function fits shown in
Galaxy clustering as a function of stellar mass

Figure 9. Projected correlation function of galaxies $w_p(\sigma)$ as a function of cumulative stellar mass $M_*$ (as labelled in each row) at redshift $z = 0.62$, computed using the Gon14 and Lac14 GALFORM models, in real (first column only) and redshift space. The VIPERS measurements of Marulli et al. (2013) are shown (black lines with error bars), whose apparent magnitude limit has been imposed on the model galaxies. For clarity, each $w_p(\sigma)$ has been divided by a reference power law $w_{\text{ref}}(\sigma)$, with parameters $r_0 = 5 h^{-1} \text{Mpc}$ and $\gamma = 2$ (horizontal line; see equation 8). Columns (a) and (b) show the default models with the true and estimated masses, respectively. The remaining columns show the results of matching the VIPERS stellar mass function (as an adjustment following the SED fitting), for (c) the default models, and (d) the models using the new SDF merger scheme.

Fig. 6 (as described in Section 4.1.2). In this way, the results shown in columns (c) and (d) of Figs 7–9 have been required to reproduce the same observationally inferred mass function fits, making use of the ordering of galaxies in stellar mass obtained through SED fitting, but differing in whether they use (c) the default merger scheme, or (d) the new scheme.

The results shown in Fig. 5 show that the changes in the model clustering predictions due to the new merger scheme, for samples selected by true stellar mass, are larger for lower stellar masses and lower redshifts. This trend is also seen for the abundance matched estimated masses when comparing columns (c) and (d) in Figs 7–9.

Reductions in the clustering amplitude lead to improved small-scale agreement with the SDSS data for both models in Figs 7(c) and 7(d). In the SDSS comparison, the Lac14 clustering also becomes stronger in the highest mass bin, resulting in improved agreement with the observational data on large scales. However, in several mass bins in Fig. 7(d) the small-scale clustering amplitude is still high with respect to the SDSS measurements, particularly for the Gon14 model. Additionally, the model clustering predictions on the largest scales are often excessive with respect to the SDSS results in Fig. 7(d), particularly so for the two lowest mass bins.

Both models experience suppressed small-scale clustering in Figs 8(c) and 8(d), which is more significant in the lower mass bin, leading to good agreement with the small-scale GAMA measurements for the Gon14 model in the lower mass bin, and for the Lac14 model in the higher mass bin. However, the Lac14 model clustering amplitude in Fig. 8(d) is now low on small scales with respect to the GAMA result in the lower mass bin. The suppression of the Gon14 clustering in the higher mass bin, due to the new merger scheme, is small compared to the excessive clustering on small scales with respect to the GAMA measurement in Fig. 8(c).

In Figs 9(c) and 9(d), suppression of the small-scale clustering tends to push the $w_p(\sigma)$ predicted by the Gon14 model closer to the VIPERS data, and further below this for the Lac14 model; except in the highest mass bin, where the Lac14 model clustering amplitude rises. However, the impact of the new merger scheme on the clustering predictions shown in Fig. 9(d) is relatively small, and the overall level of agreement with the VIPERS measurements is almost unchanged from the results shown in Fig. 9(c).

5.2.4 The need for SED fitting

We have used abundance matching of model stellar mass functions to observational measurements as a basis for comparing the clustering predictions of the different GALFORM models to each other, and to observational estimates. The abundance matching has been carried out as a final adjustment to the estimated model galaxy masses. These results are shown in columns (c) and (d) of Figs 7–9, and have
been discussed in Sections 5.2.2 and 5.2.3. It is interesting to ask how these results, which have undergone mass function matching following SED fitting, compare to those obtained by simply matching the observationally inferred mass functions using the true model stellar masses. This test examines how important the SED fitting is when making clustering predictions to compare to observations, and in particular how significantly the ordering of galaxies in stellar mass changes when carrying out SED fitting (in terms of the impact on the clustering).

In a small number of cases, the changes due to the SED fitting are insignificant when using imposed mass functions. For example, the results for the lowest mass GAMA galaxies (as shown in Figs 8c and 8d) are unperturbed by SED fitting for both the Gon14 and Lac14 models, with or without the new merger scheme. Yet for the higher mass GAMA galaxies, not carrying out SED fitting results in clustering predictions using the new merger scheme which on small scales are midway in log_{10} \( w_p(\sigma) \) between those shown in Fig. 8(d), for both GALFORM models. Similar trends are obtained in the highest VIPERS mass interval, for example, where for either merger scheme, both models yield small-scale clustering amplitudes midway between the model lines shown in Figs 9(c) and 9(d).

In general, we find that without SED fitting, the models with matched mass functions produce clustering predictions which are different to those of their counterparts based on estimated masses, particularly for \( \sigma \lesssim 1 \ h^{-1} \text{ Mpc} \).

5.2.5 Sources of uncertainty

We now discuss sources of uncertainty which are relevant to the comparisons to observational data carried out in this work.

(i) A potential complication when comparing to observational galaxy clustering measurements is the inability of spectrographs to resolve galaxies in close proximity on the sky (due to fibre collisions). This effect systematically lowers the clustering amplitude, particularly on small scales (e.g. Zehavi et al. 2002; Pollo et al. 2005; de la Torre et al. 2013).

For SDSS, spectroscopic fibres on a given plate must be separated by at least 55 arcsec, which means that at the median redshift of Li et al. (2006), two galaxies cannot be observed simultaneously within \( \sigma \lesssim 0.1 \ h^{-1} \text{ Mpc} \), although this can influence \( w_p(\sigma) \) out to larger scales. Li et al. carry out a correction to account for this, based on comparing the projected correlation function to that derived using the parent photometric catalogue. It is possible that, despite this correction, there remains some systematic bias in the observational results on the small scales shown in Fig. 7.

GAMA employs a sophisticated ‘greedy’ tiling strategy, yielding nearly spatially uniform spectroscopic completeness (Robotham et al. 2010; Driver et al. 2011). Thus the small-scale clustering measurements of Farrow et al. (in preparation) shown in Fig. 8 should be reasonably reliable (i.e. robust against fibre collisions).

A correction is made by Marulli et al. (2013) to account for spectroscopic fibre collisions in VIPERS, motivated by the Munich semi-analytic galaxy formation model of de Lucia & Blaizot (2007). Mock catalogues are constructed from this model, with and without modelling of the spectrograph selection. The relative difference between the correlation functions computed for the two sets of mock galaxies is used to impose a correction on the observed result. This model-dependent correction may produce results which are systematically different on small scales to what would be measured for a truly spectroscopically complete sample. de la Torre et al. (2013) note that for VIPERS the fraction of missing galaxy pairs becomes significant below 0.03, which at the median redshift of the data to which we compare corresponds to \( \sigma \lesssim 0.8 \ h^{-1} \text{ Mpc} \) (cf. Fig. 9).

(ii) The differences on large scales between \( w_p(\sigma) \) computed in real and redshift space for a given model in Figs 7–9 show that \( \pi_{\text{max}} = 30 \ h^{-1} \text{ Mpc} \) is not sufficiently large for the projected correlation function computed in redshift space to converge to the real space result, but is appropriate for comparison to the observational data (see Section 4.2.3). Note that we integrate the real and redshift space pair separations out to the same \( \pi_{\text{max}} \) when computing \( w_p(\sigma) \) (see Section 4.2.2). For clarity, the real space clustering is shown only in column (a) of Figs 7–9. The offsets between the real and redshift space clustering are similar for the columns where the real space lines are not shown.

(iii) Differences exist between the broad-band SED fitting procedure implemented here, and the method of Kauffmann et al. (2003) used to estimate stellar masses in the SDSS data (see Sections 3.2 and 4.1.1). However, our methodology yields stellar masses which are more appropriate to use than the true model stellar masses when comparing to these observational results.

(iv) Our SED fitting procedure uses the Calzetti et al. (2000) dust model, whereas Davidzon et al. (2013) also permit the Prevot–Bouchet model for VIPERS (see Section 3.4). However, Davidzon et al. demonstrate that using just the Calzetti et al. model has only a marginal impact on the recovered stellar mass function.

(v) As we make use of simulation snapshots in this paper (i.e. model outputs at fixed redshifts), rather than constructing light cones which cover specific redshift intervals (e.g. Merson et al. 2013), the interpretation of a sample of galaxies selected in apparent magnitude is slightly different to that for an observational survey. For example, if the model redshift is chosen to be close to the median redshift of an observational sample (as we do here), intrinsically fainter galaxies can exist in the survey sample than in the model, when the model and observations use the same faint apparent magnitude limit, due to the width of the observational redshift bin. This effect becomes important, for our purposes, for stellar masses at which the mass function becomes highly incomplete, due to the combination of redshift and apparent magnitude limits (see Section 4.1.2). In particular, this mass scale will tend to occur at higher stellar masses in a model snapshot than in a survey sample with a lower minimum redshift than the snapshot redshift. While the use of light cones could facilitate a more accurate treatment of the selections in apparent magnitude, the stellar mass ranges over which we compare the model clustering predictions to observational measurements are reasonably close to being complete in stellar mass. That is, we do not consider observational measurements where the survey selections result in a severely incomplete stellar mass function, and so the use of snapshots should not have a significant influence on the reliability of our comparisons. Light cones must be used, for example, to compare theoretical model predictions to the higher redshift GAMA clustering data of Farrow et al. (in preparation), which we do not consider here for this reason.

6 CONCLUSIONS

The stellar masses of real galaxies have to be derived from observables, typically using broad-band SED fitting. We have presented a new methodology for comparing the clustering predictions of galaxy formation models to observational data, for samples selected by stellar mass. The approach is to use estimated masses for the model galaxies, recovered using SED fitting to the model broad-band photometry. This allows us to incorporate the various systematic errors
and biases involved in the fitting procedure, which can lead to significant differences with respect to the true stellar masses. These differences cannot be properly accounted for by assuming a mean rescaling, or even a rescaling plus a scatter, between the true and estimated masses, particularly for massive galaxies with strong dust attenuation (Mitchell et al. 2013).

If stellar masses from galaxy formation models and observations can be compared in such a consistent way, the clustering of galaxies as a function of stellar mass can be used to constrain the physical processes implemented in the models, alongside traditional statistics such as the luminosity function. Our methodology demonstrates how to do such a comparison.

We have compared the clustering predictions of the Gon14 and Lac14 GALFORM models to observational measurements from different surveys at redshifts of 0.1, 0.2, and 0.6. The clustering of the model galaxies as a function of stellar mass can change significantly when moving to the estimated masses. This can be understood in terms of the transfer of galaxies between stellar mass bins when estimated masses are used, which changes the distribution of halo masses in a given bin.

Considering the estimated model masses (and in particular following abundance matching to the observed stellar mass functions), we have often found that the models predict higher small-scale clustering amplitude than is inferred observationally (for projected we have often found that the models predict higher small-scale clustering amplitude than is inferred observationally (for projected separations $\sigma \lesssim 1 \, h^{-1} \, \text{Mpc}$). This tends to be the case for both GALFORM models at the lowest redshift in our comparison (SDSS; see Fig. 7c), and for the Gon14 model at higher redshifts (comparisons to GAMA and VIPERS; see Figs 8c and 9c). In the higher redshift comparisons, the Lac14 model tends to predict similar or lower small-scale clustering amplitude than is measured using the survey data (again see Figs 8c and 9c).

The small-scale clustering is sensitive to the treatment of galaxy mergers. We have introduced a new scheme for the merging of satellites with their central galaxy, in which satellite galaxies track their associated subhalo in the dark matter simulation until this is no longer resolved. In the standard GALFORM approach, galaxies are assigned analytic merger time-scales as soon as they become satellites. The latter implementation yields higher small-scale clustering, as noted by Contreras et al. (2013), with respect to the Munich LGALAXIES semi-analytic models. The Munich models use a merger scheme which is similar to the new scheme which we have implemented in GALFORM (see Section 2.4 for a full description of the new scheme). Using this new scheme, together with estimated stellar masses abundance matched to observationally inferred stellar mass functions, generally offers improved agreement with the observational data on small scales (particularly in the SDSS and GAMA comparisons; see Figs 7c, 7d and 8c, 8d), or a similar overall level of agreement where the spread in the models encompasses the observational results (see the VIPERS comparison in Figs 9c and 9d).

Despite reductions due to the new merger scheme, the small-scale clustering often remains higher than inferred from SDSS data, particularly for the Gon14 model (see Fig. 7d). In the same comparison, the predicted clustering on the largest scales is in agreement with the SDSS measurements for only the highest masses considered. The models using the new merger scheme overall agree reasonably well with the GAMA clustering data (see Fig. 8d). However, in each mass bin, only one of the two models is in good agreement with the small-scale clustering. Comparing the same models to the VIPERS measurements in Fig. 9(d), the model that is closest to the observational constraints varies with stellar mass, and there can be significant differences between the modelled and observationally inferred clustering, on both small and large scales.

Kim et al. (2009) have shown that the galaxy clustering as a function of luminosity predicted by GALFORM is stronger on small scales than is measured from observations at low redshift, and thus argue for the inclusion in the models of mergers between satellite galaxies and their tidal disruption. Such adjustments may also improve the agreement with observational measurements of the clustering as a function of stellar mass. We have not included these enhancements here, but their potential implementation in future models, along with our new merger scheme, will produce galaxy formation predictions which take better account of the true dynamics of galaxies in dark matter haloes, and are as faithful as possible to the underlying dark matter distribution, allowing more robust constraints on the galaxy formation physics.

In summary, we have found that the model clustering predictions agree reasonably well with measurements from GAMA ($z = 0.2$). However, we have identified some relatively large discrepancies with respect to data from SDSS ($z = 0.1$) and VIPERS ($z = 0.6$), considering in particular the results shown in column (d) of Figs 7–9. The differences between the model predictions and the observations vary as a function of stellar mass, redshift, and projected separation, for each GALFORM model. The models considered in this paper have not been calibrated to reproduce any observational clustering data (or stellar mass functions). Our new methodology will enable future models to be calibrated to the clustering of galaxies as a function of stellar mass, thus providing more accurate and reliable constraints on how galaxies populate dark matter haloes.

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REFERENCES

APPENDIX A: MODEL PARAMETERS

As outlined in Section 2.3, the Gon14 and Lac14 GALFORM models are each based on the model of Lagos et al. (2012). This appendix gives details of the parameter differences between the Gon14 and Lac14 models; beyond the choice of IMF, SPS model, and default satellite merger scheme (these are described in Sections 2.3 and 2.4). Appendix A gives details of the parameter differences between the Gon14 and Lac14 models; beyond the choice of IMF, SPS model, and default satellite merger scheme (these are described in Sections 2.3 and 2.4). Table A1 gives the values of the relevant parameters in each model. We refer the reader to Cole et al. (2000), Bower et al. (2006), and Lagos et al. (2012) for further details of the model physics.
The total star formation rate of a galaxy $\psi$ is the sum of the contributions from quiescent star formation in the disc $\psi_{\text{disc}}$, and from bursts $\psi_{\text{burst}}$. The quiescent component is computed as

$$\psi_{\text{disc}} = \nu_{\text{SF}} f_{\text{mol}} M_{\text{cold, disc}},$$  \hspace{1cm} (A1)

where $M_{\text{cold, disc}}$ is the mass of cold gas in the disc, a fraction $f_{\text{mol}}$ of which is molecular. $\nu_{\text{SF}}$ is the inverse of the star formation time-scale for molecular gas (see Table A1).

The rate $\dot{M}_{\text{reheat}}$ at which cold gas is reheated by supernova feedback is given by

$$\dot{M}_{\text{reheat}} = \psi \left( \frac{V_{\text{circ}}}{V_{\text{hot}}} \right)^{-\alpha_{\text{hot}}},$$  \hspace{1cm} (A2)

where $V_{\text{circ}}$ is the circular velocity of the galaxy, at the half-mass radius. $V_{\text{hot}}$ is a parameter with velocity units (see Table A1), and $\alpha_{\text{hot}} = 3.2$ for both models. Gas heated by supernova feedback is assumed to be ejected beyond the virial radius, and then reincorporated into the hot halo gas reservoir at a rate proportional to $\alpha_{\text{reheat}}$ (see Table A1).

The onset of AGN feedback in GALFORM requires the cooling time $t_{\text{cool}}$ and free-fall time $t_{\text{ff}}$ of the gas to satisfy

$$t_{\text{cool}} > t_{\text{ff}}/\alpha_{\text{cool}},$$  \hspace{1cm} (A3)

that is, the halo is assumed to be in a state of quasi-hydrostatic cooling. This criterion is controlled by the dimensionless parameter $\alpha_{\text{cool}}$ (see Table A1). Additionally, the cooling rate of the gas must be less than a fraction $\epsilon_{\text{SMBH}}$ of the Eddington luminosity (see Table A1). If both conditions are satisfied, AGN feedback is assumed to suppress the cooling of the halo gas.

The orbital energy of merging galaxies is assumed to be proportional to $f_{\text{orbit}}$, and is used to compute the size of the spheroid formed in the merger. $f_{\text{burst}}$ is the minimum galaxy mass (stars plus cold gas) ratio in a minor merger required to trigger bursts of star formation (major mergers occur for mass ratios above 0.3). Bursts are suppressed in minor mergers if the gas fraction in the disc of the main galaxy is less than $f_{\text{gasburst}}$. The star formation time-scale for a burst is given either by $f_{\text{dyn}}$ times the bulge dynamical time, or by the parameter $\tau_{\text{min}}$, whichever is larger (see Table A1).

Disc instabilities, with associated bursts of star formation, are triggered when

$$\frac{V_{\text{circ}}}{(GM_{\text{disc}}/r_{\text{disc}})^{2/3}} < \epsilon_{\text{disc}},$$ \hspace{1cm} (A4)

where $M_{\text{disc}}$ and $r_{\text{disc}}$ are the mass of the disc and its half-mass radius, respectively, and $\epsilon_{\text{disc}}$ is a parameter (see Table A1).

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### Table A1

Parameters which differ between the Gon14 and Lac14 models. A description of each parameter is given in the main text.

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<th>Parameter</th>
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<th>Lac14</th>
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<td>$f_{\text{dyn}}$</td>
<td>10</td>
<td>20</td>
</tr>
<tr>
<td>$\tau_{\text{min}}$ [Gyr]</td>
<td>0.05</td>
<td>0.1</td>
</tr>
<tr>
<td>$\epsilon_{\text{disc}}$</td>
<td>0.8</td>
<td>0.9</td>
</tr>
</tbody>
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